

# Supplement 4

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5/20/2020

```
library(knitr)
opts_chunk$set(tidy.opts=list(width.cutoff=60),tidy=TRUE)
```

## Weight consumption (WC)

```
WC<-read_excel("Table S1.xls") #import Table S1
WC_df<-as.data.frame(WC)
WC_df$species <- as.factor(WC_df$species)
#ANOVA using linear model using function "anova"
WC.lme.base1<-lm(wc~species, data=WC_df)
WC_lme_base1_aov<-aov(WC.lme.base1)
```

## WC:fitting a random effect

```
#Fitting a random effect
WC_lme_ran<-lme(wc ~ species, random=~ 1|trial, data=WC_df, method = "REML")
```

## WC:comparing models

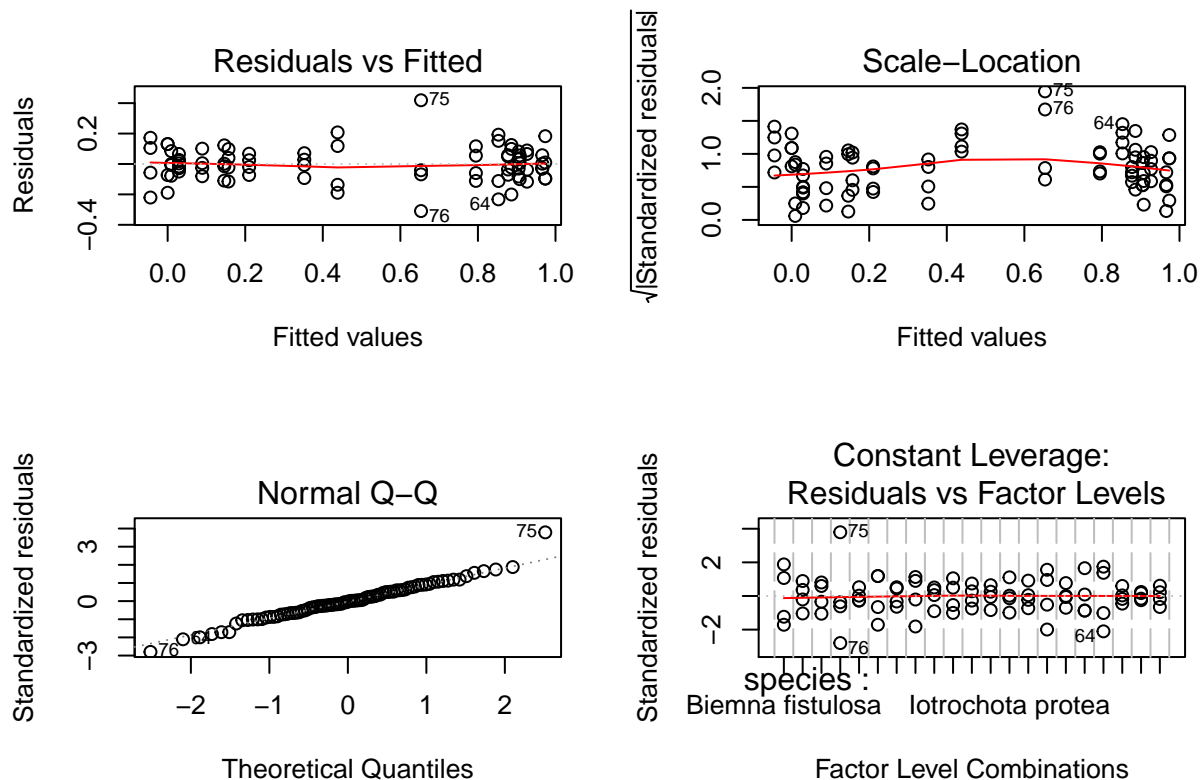
```
#compare model
testmodel<-anova(WC_lme_ran,
  WC.lme.base1)
testmodel
```

| ## | Model        | df | AIC | BIC       | logLik   | Test     | L.Ratio | p-value      |
|----|--------------|----|-----|-----------|----------|----------|---------|--------------|
| ## | WC_lme_ran   | 1  | 23  | -5.458549 | 43.83355 | 25.72927 |         |              |
| ## | WC.lme.base1 | 2  | 22  | -7.458549 | 39.69042 | 25.72927 | 1 vs 2  | 3.055028e-09 |

```
#best model is WC.lme.base1
```

## WC:Test Homogeneity of variance

```
layout(matrix(c(1,2,3,4),2,2))
plot(WC_lme_base1_aov) #Several outliers are spotted.
```



```
##WC:Test HOV
```

```
with(WC_df,leveneTest(wc~species)) #Does not pass HOV
```

```
## Levene's Test for Homogeneity of Variance (center = median)
##      Df F value Pr(>F)
## group 20  2.0963 0.01373 *
##      63
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

WC:##remove only row 64 (Mycale parishi\_trial4) from MC\_df

```
WC_df_no_64<-WC_df[-c(64), ]
WC_df_no_64.lme.base1<-lm(wc~species, data=WC_df_no_64)
WC_df_no_64.lme.base1_aov<-aov(WC_df_no_64.lme.base1)
summary(WC_df_no_64.lme.base1_aov)
```

```
##      Df Sum Sq Mean Sq F value Pr(>F)
```

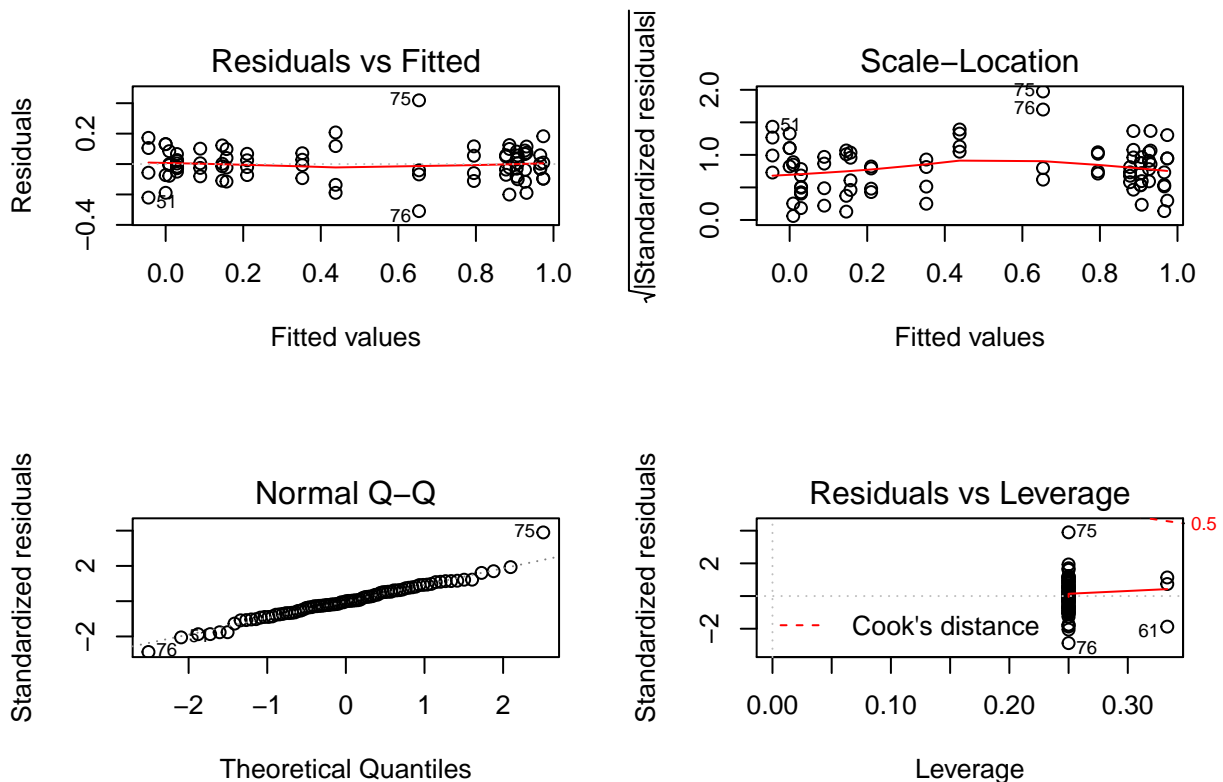
```
## species      20 12.960  0.6480  42.09 <2e-16 ***
## Residuals   62  0.955  0.0154
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
with(WC_df_no_64,leveneTest(wc~species))
```

```
## Levene's Test for Homogeneity of Variance (center = median)
##      Df F value Pr(>F)
## group 20  1.5632 0.09241 .
##      62
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

WC: #visually look for HOV

```
#HOV assumption visual
layout(matrix(c(1,2,3,4),2,2))
plot(WC_df_no_64.lme.base1)
```



```
#Lot more equality of variance of the residulas
#qqplot looks better and variance seems alot more stable
#Lets test levene's test to see if HOV passes
```

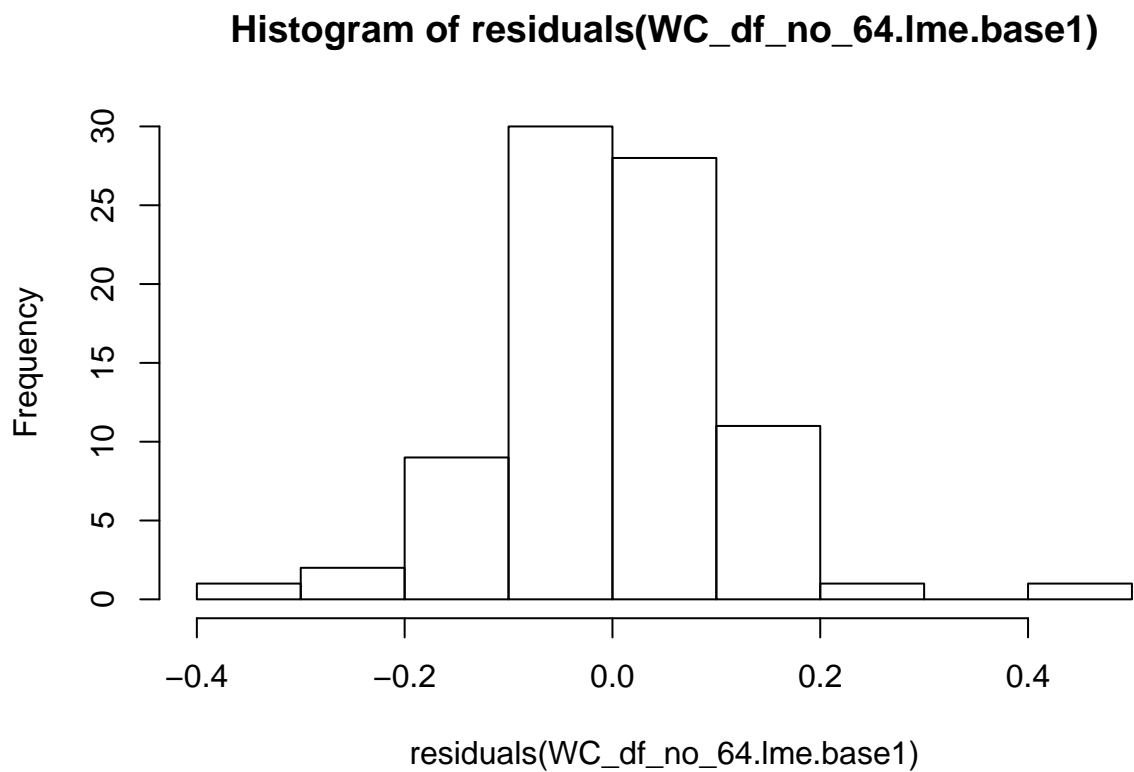
## WC:Levene's test HOV

```
#HOV assumption-Levene's test
with(WC_df_no_64,leveneTest(wc~species))

## Levene's Test for Homogeneity of Variance (center = median)
##      Df F value Pr(>F)
## group 20  1.5632 0.09241 .
##      62
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

## WC:histogram

```
hist(residuals(WC_df_no_64.lme.base1))
```



```
##WC:Anderson Darlington Normality test
```

```
#Anderson Darlington
ad.test(residuals(WC_df_no_64.lme.base1))
```

```
##
```

```
## Anderson-Darling normality test
##
## data: residuals(WC_df_no_64.lme.base1)
## A = 0.42987, p-value = 0.3016
```

### WC:Cramer-von Mises normality test

```
#Cramer-von Mises
cvm.test(residuals(WC_df_no_64.lme.base1))
```

```
##
## Cramer-von Mises normality test
##
## data: residuals(WC_df_no_64.lme.base1)
## W = 0.04989, p-value = 0.5077
```

### WC:Lilliefors (Kolmogorov-Smirnov) normality test

```
lillie.test(residuals(WC_df_no_64.lme.base1))
```

```
##
## Lilliefors (Kolmogorov-Smirnov) normality test
##
## data: residuals(WC_df_no_64.lme.base1)
## D = 0.064016, p-value = 0.5486
```

### WC:Pearson chi-square normality test

```
pearson.test(residuals(WC_df_no_64.lme.base1))
```

```
##
## Pearson chi-square normality test
##
## data: residuals(WC_df_no_64.lme.base1)
## P = 4.4699, p-value = 0.8779
```

```
pearson.test(residuals(WC_anova2)) Pearson chi-square normality test
data: residuals(WC_anova2) P = 2.2857, p-value = 0.9861
```

### WC:ANOVA results of model (WC\_df\_no\_64.lme.base1\_aov)

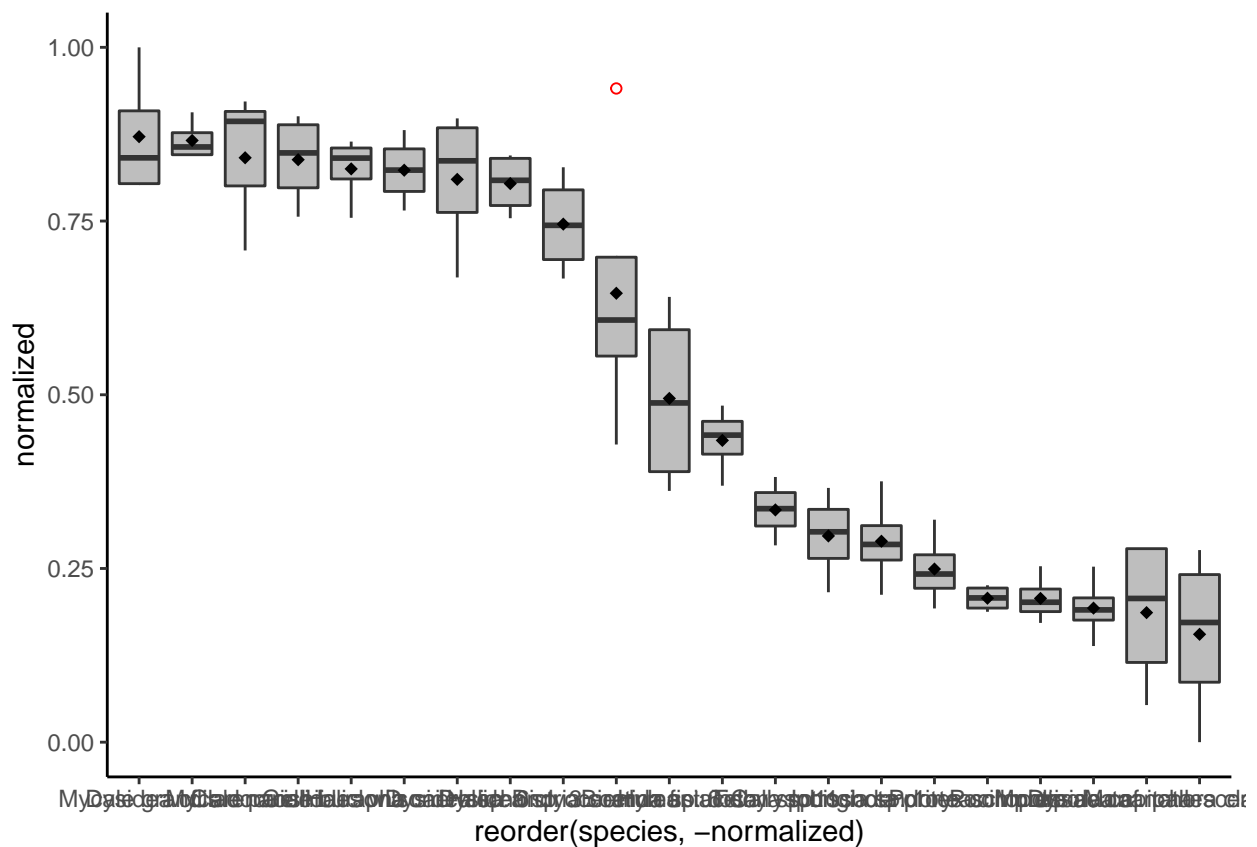
```
summary(WC_df_no_64.lme.base1_aov)
```

```
##           Df Sum Sq Mean Sq F value Pr(>F)
## species    20 12.960  0.6480  42.09 <2e-16 ***
## Residuals   62  0.955  0.0154
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

## WC:Tukey's post hoc analysis

```
#post-hoc analysis
WC_Tukey<-TukeyHSD(WC_df_no_64.lme.base1_aov)
WC_Tukey <- TukeyHSD(WC_df_no_64.lme.base1_aov)
# the [1:1] locates the part of the output to be exported
WC_Tukey_df<-as.data.frame(WC_Tukey[1:1])
# convert row name into first column
WC_Tukey_df<- cbind(rownames(WC_Tukey_df), data.frame(WC_Tukey_df, row.names=NULL))
write_csv(WC_Tukey_df, "WC_Tukey_df.csv")
#plot
Fig.2a <- ggplot(WC_df_no_64, aes(x=reorder(species,-normalized), y=normalized)) +
  stat_summary(fun=mean, geom="point", shape=23, size=1)+
  geom_boxplot(outlier.colour = "red", outlier.shape = 1, fill="grey")+
  stat_summary(fun=mean, geom="point", shape=18, size=2, color="black", fill="black")+
  theme_classic()
```

Fig. 2a



```
##WC:Heatmap plot to generate Fig. 2b
```

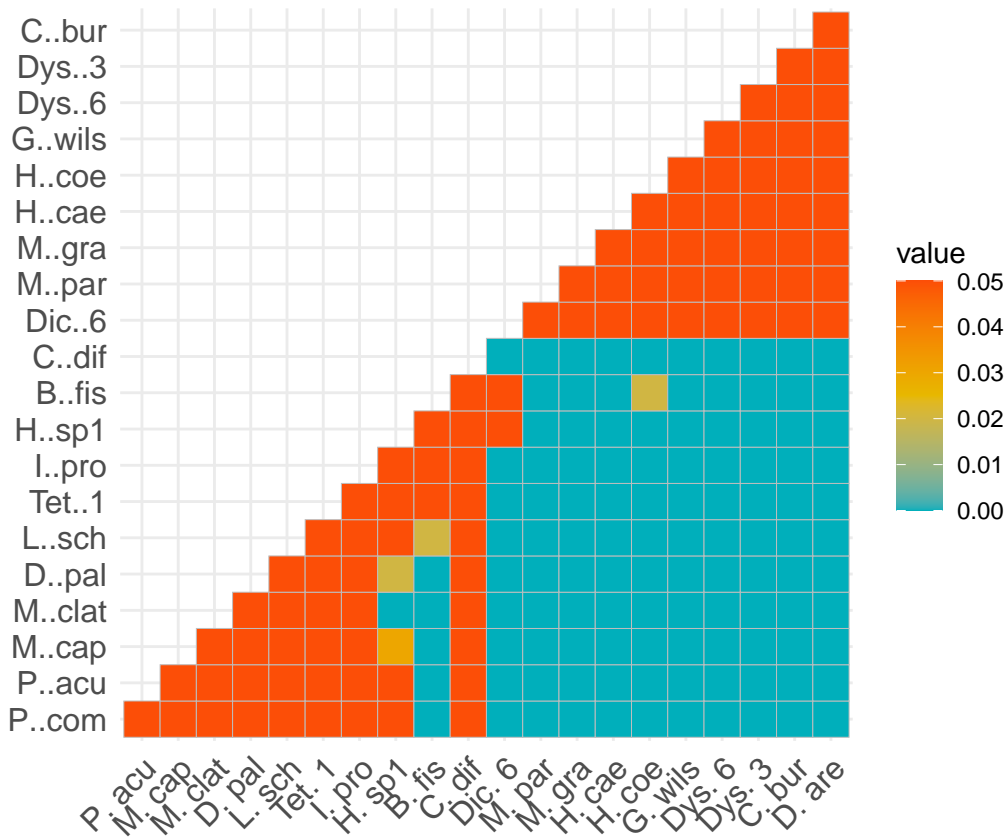
```
#heatmap to look at significant Tukey results and generate Fig. 2b. We reorganized  
#significant p values from pairwise comparisons in Table S5 into a correlation matrix  
#(TableS10).  
#TableS10 is not provided.  
hp<-read_excel("TableS10.xls")
```

```
## New names:  
## * `` -> ...1
```

```
hp_df<-data.frame(hp, row.names=1)  
hp_df[hp_df > .05] <- 0.05 #converting all values >0.05 to 0.05  
hp_df.mat<-as.matrix(hp_df)  
view(hp_df)  
Fig.2b<-ggcorrplot(hp_df.mat, type="lower", hc.order=TRUE, method = "square") +  
  scale_fill_gradient2(limit = c(0,0.05), low = "#00AFBB", high = "#FC4E07",  
    mid = "#E7B800", midpoint = 0.025)
```

```
## Scale for 'fill' is already present. Adding another scale for 'fill', which  
## will replace the existing scale.
```

Fig.2b



## SAC data analysis

```
# references
#https://ncss-wpengine.netdna-ssl.com/wp-content/themes/ncss/pdf/
#Procedures/NCSS/Mixed_Models-Repeated_Measures.pdf;
#https://rcompanion.org/handbook/I_09.html; https://ademos.people.uic.edu/Chapter18.html
##citation:Mangiafico, S.S. 2016. Summary and Analysis of Extension Program
##Evaluation in R, version 1.18.1. rcompanion.org/handbook/.
#(Pdf version rcompanion.org/documents/RHandbookProgramEvaluation.pdf.)

SAC<-read_excel("Table S2.xls")# Importing and formatting data excel sheets to R
#statistical analysis
SAC_df<-as.data.frame(SAC)
SAC_df$correction<-as.numeric(SAC$correction)
SAC_df$hours <- factor(SAC$hours)#making hours a factor
```

## SAC: fitting a linear model with fixed effects

```
#For the corAR1 structure, a value for the first order
#correlation can be specified. In this case,
#the value of 0.429 is found using the ACF function
#determining autocorrelation in residuals

#Fit linear model using generalized least squares
SAC.model.gls = gls(correction ~ species + hours + species*hours,
                    data=SAC_df)

SAC.model.gls.ACF<-ACF(SAC.model.gls)
view(SAC.model.gls.ACF)
```

## SAC:determining autocorrelation in residuals

## SAC:linear mixed effects base model

```
SAC.model.lme = lme(correction ~ species + hours + species*hours,
                    random = ~1|trial,
                    data=SAC_df)
SAC.model.lme.ACF<-ACF(SAC.model.lme)
SAC.model.lme.ACF
```

```
##      lag      ACF
## 1     0 1.000000000
## 2     1 0.5549906734
## 3     2 0.2866825429
## 4     3 0.1197080976
## 5     4 -0.0974173112
## 6     5 -0.1530004105
## 7     6 -0.2003259008
```



```
## 8 7 -0.1748970127
## 9 8 -0.1465801341
## 10 9 -0.0762675327
## 11 10 -0.0227106137
## 12 11 0.0217988360
## 13 12 0.0495616115
## 14 13 0.0003156655
## 15 14 -0.0397561815
## 16 15 -0.1433857252
## 17 16 -0.0827107547
## 18 17 -0.0441271181
```

```
# Value 0.5549906734 is the value for the first order correlation of model SAC.model.lme
```

## SAC:fixed effects

```
#fixed effects model
SAC.model.fixed = gls(correction ~ species + hours + species*hours,
                      data=SAC_df,
                      method="REML")
```

## SAC:linear mixed effects base mode with regressive lag structure

```
#linear mixed effects base model with autoregressive lag 1 structure
SAC.model.coAR1.rand = lme(correction ~ species + hours + species*hours,
                          random = ~1|trial,
                          correlation = corAR1(form = ~ 1|trial, value = 0.5550),
                          data=SAC_df,
                          method="REML")
```

## SAC:comparing models

```
#test random effects
testmodel<-anova(SAC.model.fixed, SAC.model.coAR1.rand, SAC.model.lme)
testmodel
```

```
##           Model df      AIC      BIC    logLik   Test  L.Ratio
## SAC.model.fixed      1 55 178.6604 348.4782 -34.33018
## SAC.model.coAR1.rand  2 57 121.9944 297.9874 -3.99721 1 vs 2 60.66594
## SAC.model.lme        3 56 179.6868 352.5922 -33.84341 2 vs 3 59.69239
##           p-value
## SAC.model.fixed
## SAC.model.coAR1.rand <.0001
## SAC.model.lme       <.0001
```

```
#random effects seems to be better fit (AIC=121.99)
```

## SAC:ANOVA analysis of the model

```
SAC.model.coAR1.rand_anova<-anova(SAC.model.coAR1.rand, type = "marginal")  
SAC.model.coAR1.rand_anova
```

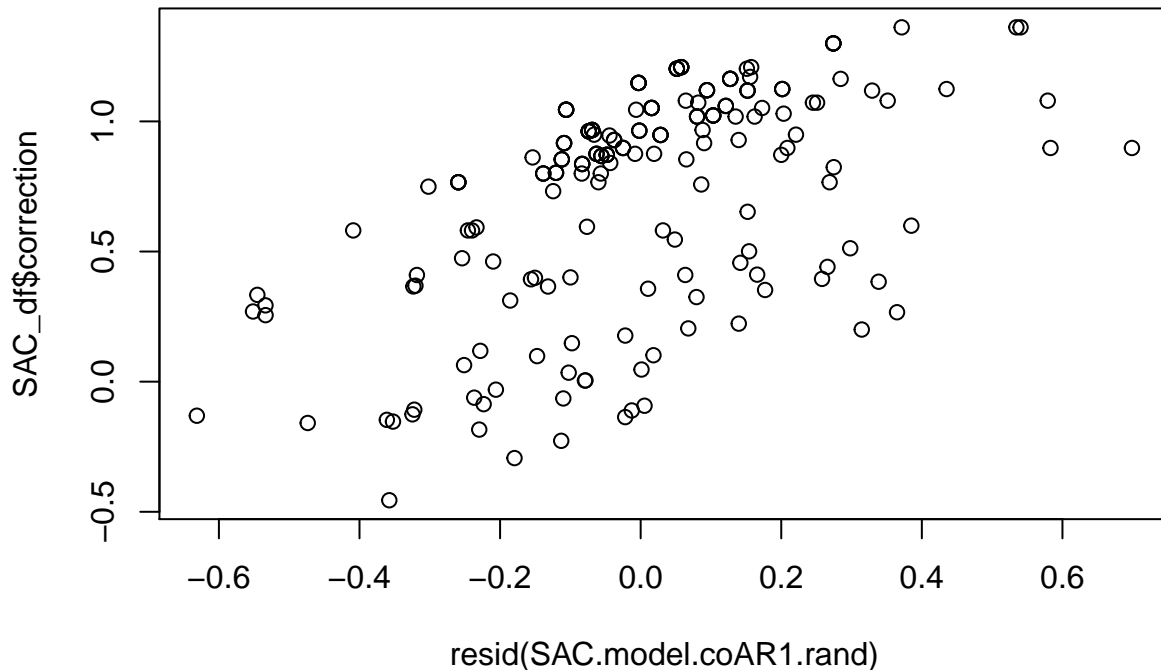
```
##           numDF denDF  F-value p-value  
## (Intercept)      1  159  0.497054  0.4818  
## species          8  159 11.596270 <.0001  
## hours           5  159  9.499972 <.0001  
## species:hours   40  159  2.563586 <.0001
```

## SAC:Tukey's post hoc analysis to determine pairwise differences

```
SAC.model.coAR1.rand.tukey <- emmeans(SAC.model.coAR1.rand,  
                                     list(pairwise ~ species*hours),  
                                     adjust = "tukey", cld = FALSE)  
SAC.model.coAR1.rand.tukey_df <-  
  as.data.frame(SAC.model.coAR1.rand.tukey$`pairwise differences of species`)  
write_csv(SAC.model.coAR1.rand.tukey_df, "SAC.model.coAR1.rand.tukey_df.csv")  
#used to Table S2 and determine letter display which show significant  
#differences between species Tukey's pairwise comparisons in Fig. 3.
```

## SAC:Testing Linearity assumption

```
SAC.model.rand.lin<-plot(resid(SAC.model.coAR1.rand), SAC_df$correction)
```



*#Seems to be randomly distributed points*

### SAC:TEsting homogeneity of variace

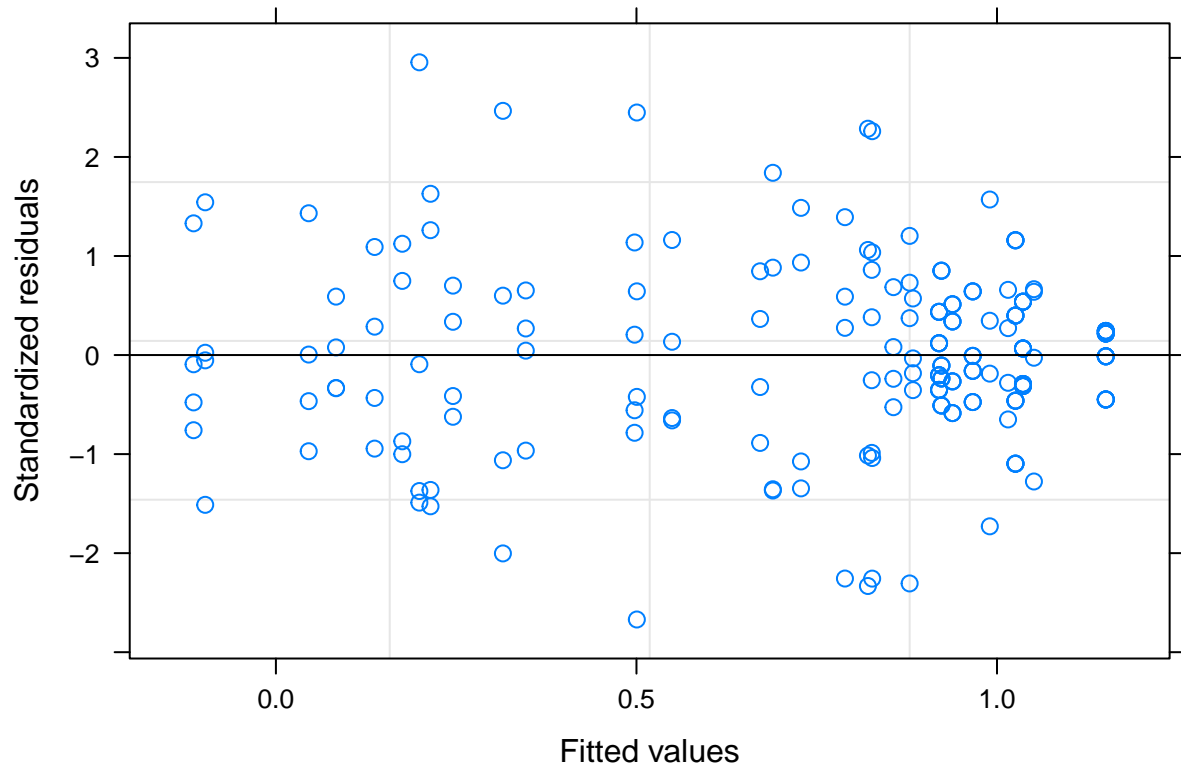
```
#extracts the residuals and places them in a new column in our original data table
SAC_df$mod_resid<- residuals(SAC.model.coAR1.rand)
#creates a new column with the absolute value of the residuals
SAC_df$Abs_mod_resid <-abs(SAC_df$mod_resid)
#squares the absolute values of the residuals to provide the more robust estimate
SAC_df$mod_resid2 <- SAC_df$mod_resid^2
#ANOVA of the squared residuals
Levene.SAC.model.rand <- lm(mod_resid2 ~ trial, data=SAC_df)
anova(Levene.SAC.model.rand)
```

```
## Analysis of Variance Table
##
## Response: mod_resid2
##          Df Sum Sq Mean Sq F value Pr(>F)
## trial      1 0.00001 0.0000130  0.0023 0.9621
## Residuals 214 1.23382 0.0057655
```

```
# variance of the residuals is equal because p value > 0.05.
```

## SAC:Visual solution

```
plot(SAC.model.coAR1.rand)
```

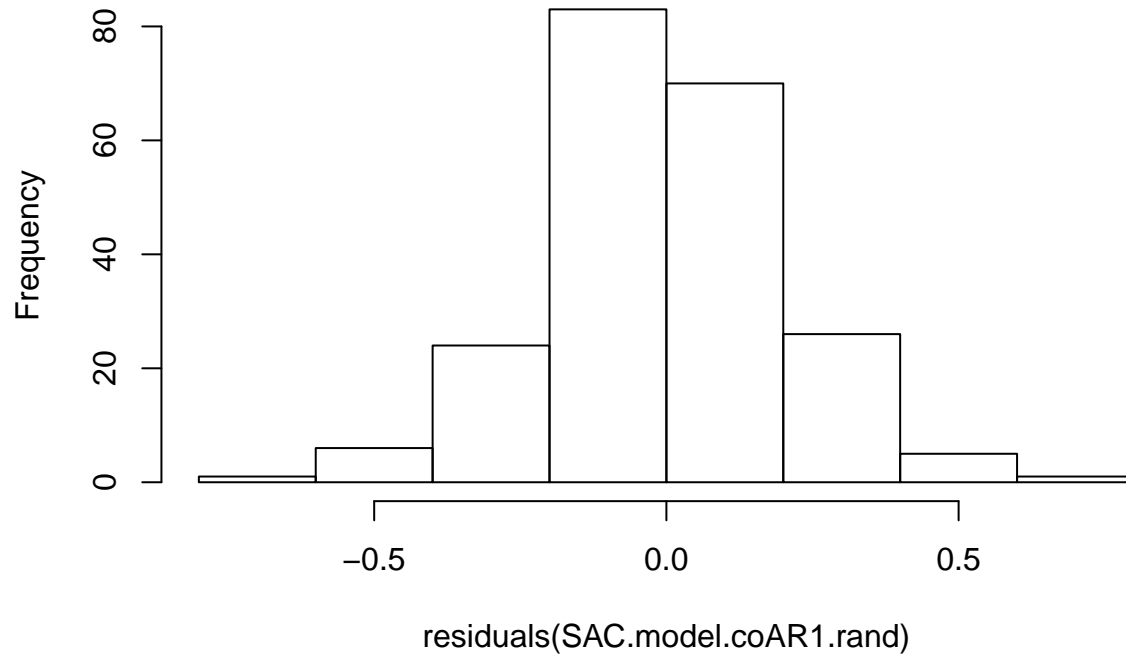


```
#Seems to show even spread around the centered line
```

## SAC:Histogram of model

```
hist(residuals(SAC.model.coAR1.rand))
```

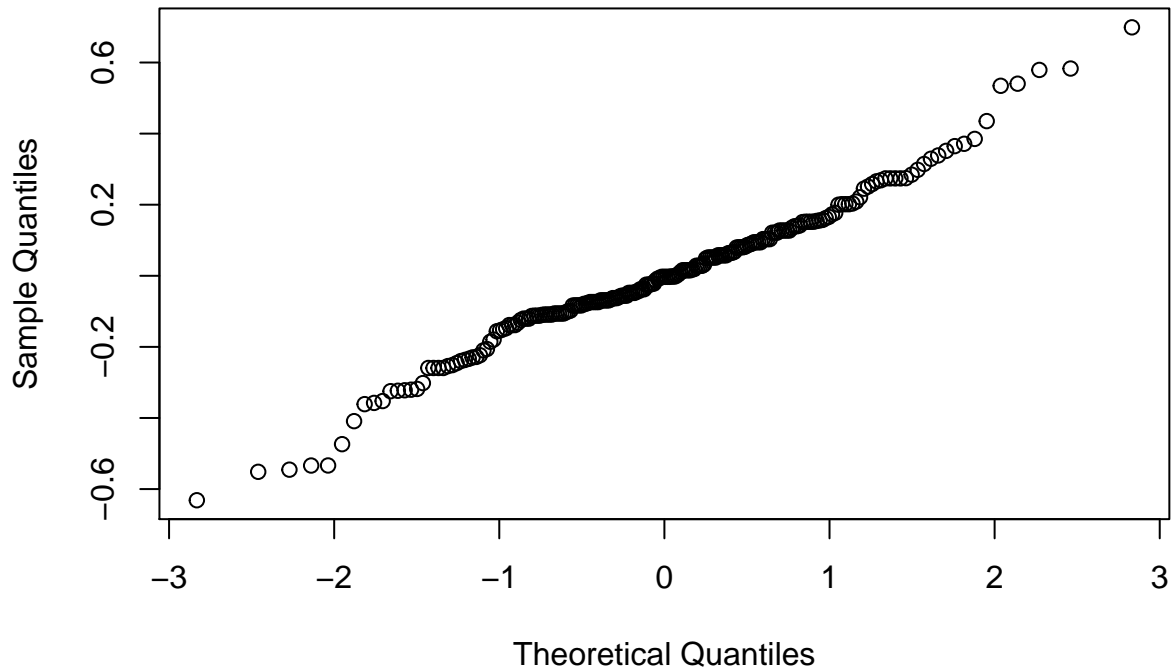
## Histogram of residuals(SAC.model.coAR1.rand)



##SAC:Assumption 3: Residuals of the model are normally distributed

```
qqnorm(resid(SAC.model.coAR1.rand, Main="qqplot"))
```

## Normal Q-Q Plot

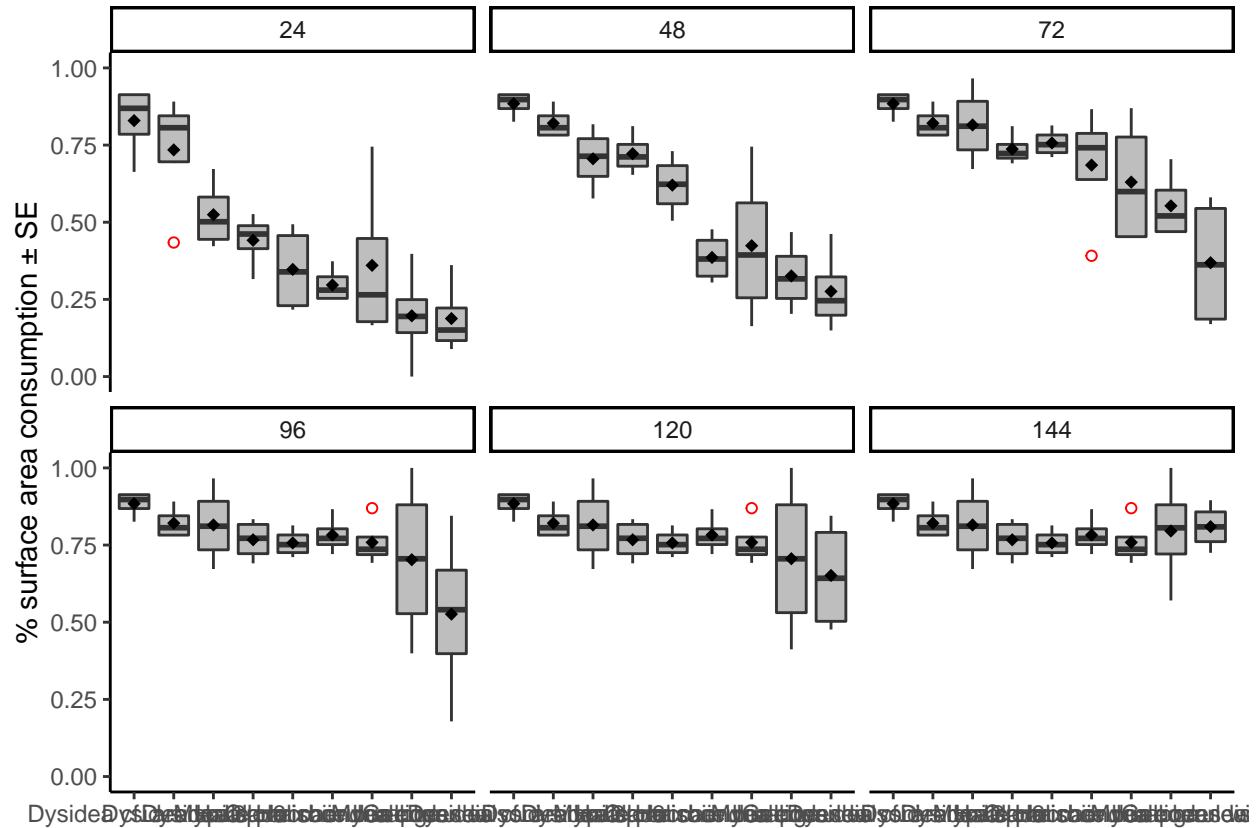


*#identifies outliers that may be exerting undue influence on the model.  
#In this case there don't seem to be any outliers.*

SAC: Subsetting the data to only include the first three time points for graphing purposes

```
Fig.3a<-ggplot(SAC_df, aes(x=reorder(species, -normalized), y=normalized)) +  
  theme_classic() + theme(legend.title = element_blank(), axis.text.x =  
    element_text(angle = 45, hjust = 1, size = 6)) +  
  geom_boxplot(outlier.colour = "red", outlier.shape = 1, fill="grey") +  
  stat_summary(fun=mean, geom="point", shape=18, size=2, color="black", fill="black") +  
  theme_classic() +  
  facet_wrap(~hours) +  
  labs(y = "% surface area consumption ± SE", x = element_blank())
```

Fig.3a



### SAC:Heatmap for Fig. 3

### SAC:heatmaps at 24 hour Fig. 3

```
#heatmap to look at significant Tukey results and generate Fig. 2b.
#We reorganized significant p values from pairwise comparisons in Table 7
#into a correlation matrix (hm24,hm48,hm72,hm96).These tables are not provided
hp24<-read_excel("hm24.xls")
```

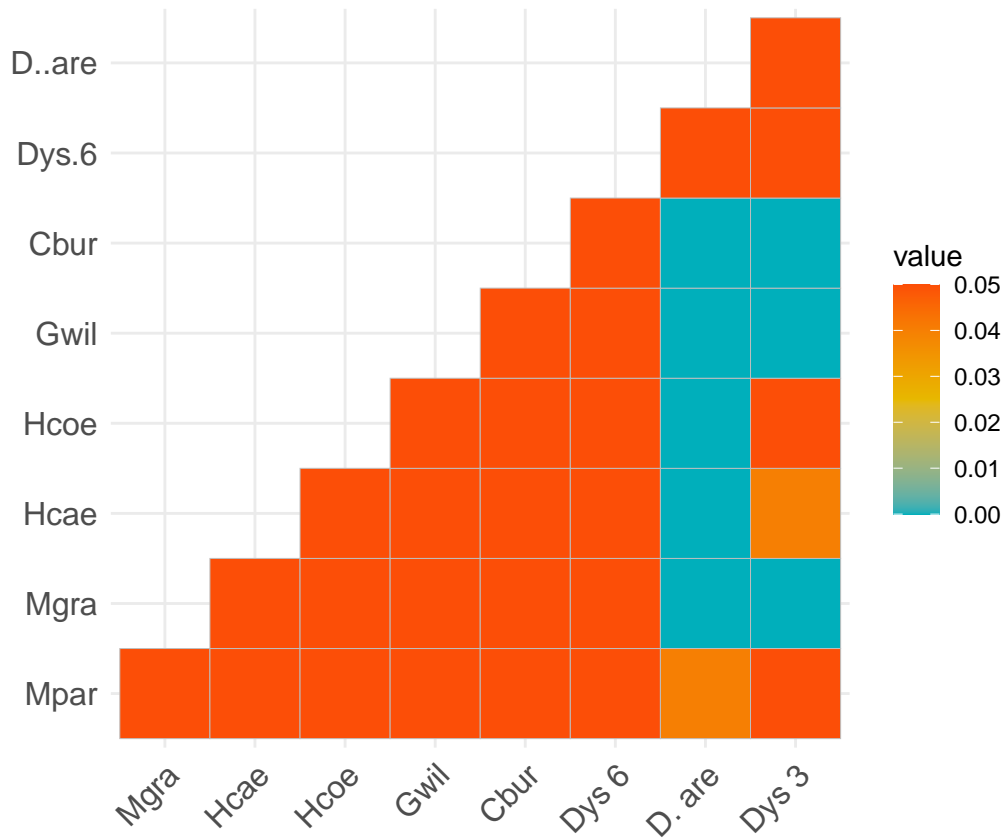
```
## New names:
## * `` -> ...1
```

```
hp_df24<-data.frame(hp24, row.names=1)
hp_df24[hp_df24 > .05] <- 0.05 #converting all values >0.05 to 0.05
hp_df24.mat<-as.matrix(hp_df24)
view(hp_df24)
```

```
Fig.3hm_24<-ggcorrplot(hp_df24.mat, type="lower", hc.order=TRUE, method = "square") +
  scale_fill_gradient2(limit = c(0,0.05), low = "#00AFBB", high = "#FC4E07" ,
    mid = "#E7B800", midpoint = 0.025)
```

```
## Scale for 'fill' is already present. Adding another scale for 'fill', which
## will replace the existing scale.
```

Fig.3hm\_24



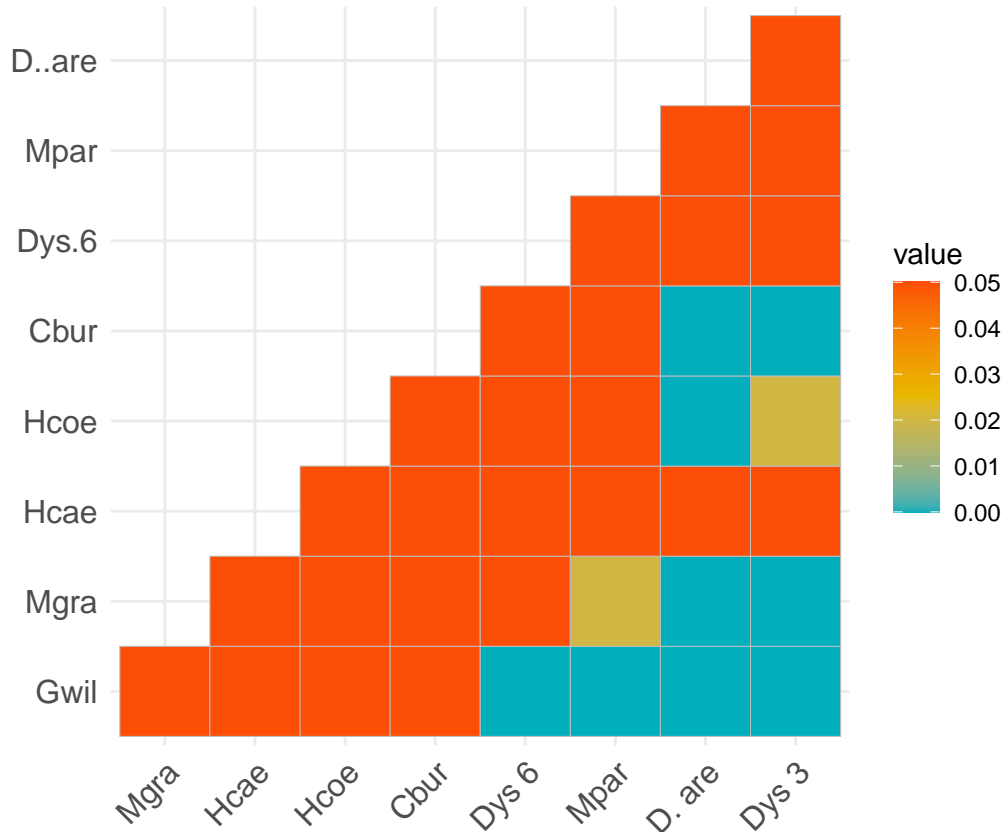
##SAC:heatmaps at 48 hour Fig. 3

```
hp48<-read_excel("hm48.xls")
hp_df48<-data.frame(hp48, row.names=1)
hp_df48[hp_df48 > .05] <- 0.05 #converting all values >0.05 to 0.05
hp_df48.mat<-as.matrix(hp_df48)
Fig.3hm_48<-ggcorrplot(hp_df48.mat, type="lower",
  hc.order=TRUE, method = "square") +
  scale_fill_gradient2(limit = c(0,0.05), low = "#00AFBB",
    high = "#FC4E07" , mid = "#E7B800", midpoint = 0.025)
```

## Scale for 'fill' is already present. Adding another scale for 'fill', which  
## will replace the existing scale.

Fig.3hm\_48



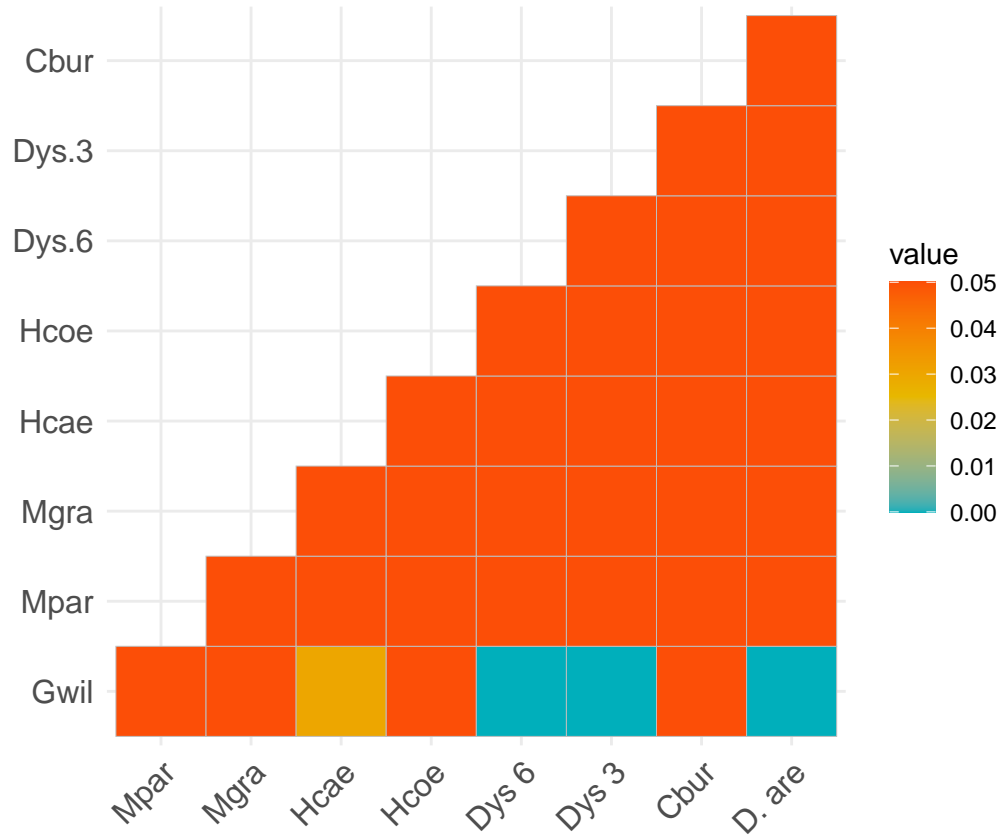


##SAC:heatmaps at 72 hour Fig. 3

```
hp72<-read_excel("hm72.xls")
hp_df72<-data.frame(hp72, row.names=1)
hp_df72[hp_df72 > .05] <- 0.05 #converting all values >0.05 to 0.05
hp_df72.mat<-as.matrix(hp_df72)
Fig.3hm_72<-ggcorrplot(hp_df72.mat, type="lower", hc.order=TRUE,
                        method = "square") +
  scale_fill_gradient2(limit = c(0,0.05),
                       low = "#00AFBB", high = "#FC4E07" , mid = "#E7B800",
                       midpoint = 0.025)
```

## Scale for 'fill' is already present. Adding another scale for 'fill', which  
## will replace the existing scale.

Fig.3hm\_72

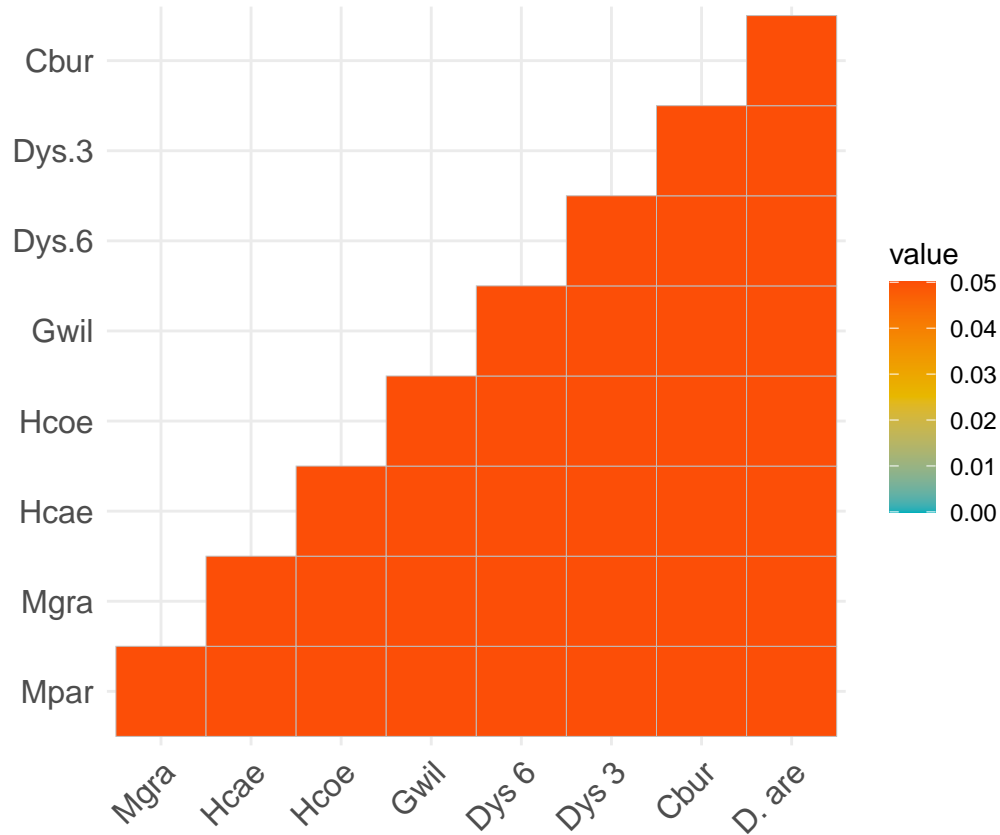


##SAC:heatmaps at 96 hour Fig. 3

```
hp96<-read_excel("hm96.xls")
hp_df96<-data.frame(hp96, row.names=1)
hp_df96[hp_df96 > .05] <- 0.05 #converting all values >0.05 to 0.05
hp_df96.mat<-as.matrix(hp_df96)
Fig.3hm_96<-ggcorrplot(hp_df96.mat, type="lower", hc.order=TRUE,
                        method = "square") + scale_fill_gradient2(limit = c(0,0.05),
                                                                    low = "#00AFBB",
                                                                    high = "#FC4E07" ,
                                                                    mid = "#E7B800",
                                                                    midpoint = 0.025)
```

## Scale for 'fill' is already present. Adding another scale for 'fill', which  
## will replace the existing scale.

Fig.3hm\_96



Note that the `echo = FALSE` parameter was added to the code chunk to prevent printing of the R code that generated the plot.

## palatability assay of crude extracts from prey species (Fig.4)

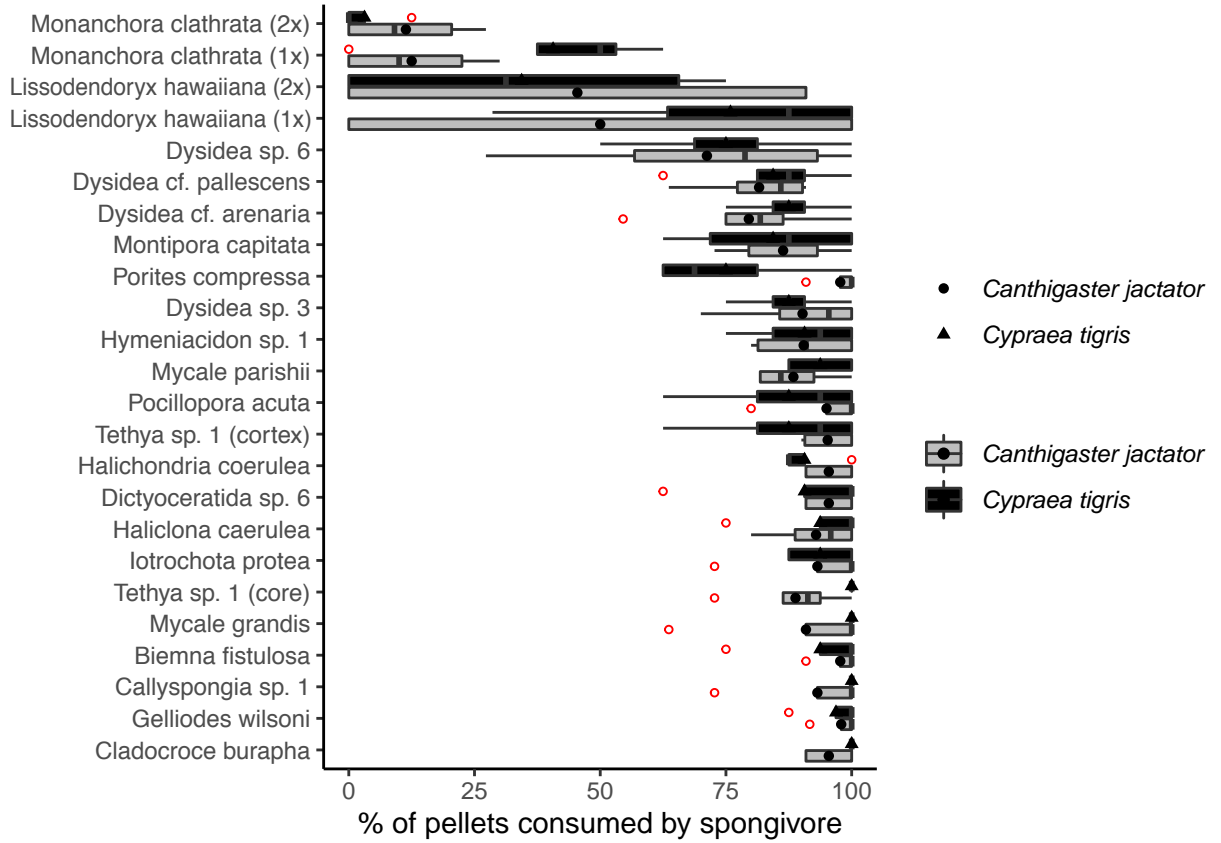
```

###Chemical palatability script for Fig. 4###
chem <-read_excel("Table S9.xls")

Fig.4<-ggplot(chem, aes(x=reorder(Species, -PercentAccept), y=PercentAccept,
                             fill=Spongivore)) +
  geom_boxplot(outlier.colour = "red", outlier.shape = 1, outlier.size = 1) +
  scale_fill_manual("legend", values=c("Canthigaster jactator" = "grey",
                                       "Cypraea tigris" = "black")) + theme_classic() +
  theme(legend.title = element_blank(),
        legend.text = element_text(face = "italic"))+
  labs(y = "% of pellets consumed by spongivore", x = element_blank())+
  stat_summary(fun=mean, geom = "point",
              size=1.5, aes(shape = Spongivore),
              position = position_dodge(width = .75))+

  rotate()
Fig.4

```



##END##