ONLINE SUPPLEMENTARY MATERIAL

Text S1. Sandeel data sources

Length at date data for the lesser sandeel (*Ammodytes marinus*) were compiled from a variety of sources (Table S1). The largest and most consistent data sources were (1) age 0 sandeel data collected weekly between 2000 and 2016 at the MSS coastal ecosystem monitoring site (using an ichtyoplankton1m ringnet in Table S1, referred to as "ringnet" in the model) on the Scottish east coast at Stonehaven (SCObs; http://data.marine.gov.scot/dataset/ scottish-coastal-observatory-stonehaven-site), (2) age 0 sandeel data measured from puffin diet on the Isle of May during the breeding season between 2000 and 2016 (full details of methods in Wanless et al. 2018, referred to as "puffins" in the model) and (3) age 0 sandeel data originating from the dedicated annual winter sandeel dredge survey targeting overwintering sandeel in the Firth of Forth (referred to as "dredge" in the model). In addition to these 3 main sources, data on age 0 sandeel caught from plankton samplers (Methot net: referred to as "MT" in the model, Gulf III sampler: referred to as "G3" in the model, pelagic (pelagic net: referred to as "pelagic" in the model and International Young Gadoid Pelagic Trawl: referred to as "PT154" in the model and demersal trawls (referred to as "demersal" in the model) and benthic samplers (modified sandeel dredge referred to as "SDG" for the summer sandeel surveys and "dredge" for the winter dredge surveys) were included (Table S1).

Sandeel abundance indices were derived from the dedicated annual winter (November/ December) sandeel dredge survey in years 2000 to 2003 and 2008 to 2016 (ICES 2021). The age 0 index from this winter survey is highly correlated with that estimated from pelagic trawls in June surveys (Régnier et al. 2017). Year-class estimates of age 0 and age 1 from winter surveys are also strongly correlated (ICES 2021). Due to a gap in annual surveys, sandeel abundance indices (age 0 and age 1) were missing for years 2004, 2005, 2006 and 2007. For age 0 abundances, missing values were imputed using the weighted average between predictions of two models. The first model described the relationship between age 0 sandeel abundance and the average mass of age 0 sandeels in puffin loads during chick rearing extracted from Wanless et al. (2018). The relationship was modelled using a linear model with logtransformed age 0 sandeel mass and log-transformed age 0 abundance index and provided a good fit (F_{1,10} = 17,09, R² = 0.63). The second model was a GLM describing variations in the age 0 sandeel abundance index with the overlap index between sandeel hatching and egg production in their copepod prey. The overlap index was extracted from Régnier et al. (2019) and the GLM provided a good fit (pseudo $R^2 = 0.74$). The respective R^2 and pseudo R^2 were use as weights to produce the weighted average between the two model predictions. Missing age 1 abundances were derived from the linear relationship between log-transformed abundance at age 1 at year *y* and log transformed abundance of age 0 at year *y-1* using annual abundance indices estimated in the Firth of Forth during periods 1999-2003 and 2008-2020 $(F_{1,12} = 14.63, R² = 0.55).$

Table S1. Length at date data (number of individuals) classified by data source.

Text S2. Seabird data sources

Median annual laying dates came from daily observations of pairs of guillemots and razorbills and observations every five days for kittiwakes and seven days for shags breeding in study plots scattered through the colony (Newell et al. 2015, Table S2). Laying dates of puffins could not be recorded directly because this species is sensitive to disturbance during incubation (Harris & Wanless 2011). Instead laying dates were back-calculated from the annual first hatching dates using a 42 days for the incubation period (Harris & Wanless 2011). Annual values for hatching dates were calculated as the annual date from standardised monitoring of puffin (first), guillemot (median) and razorbill (median). Hatching dates of kittiwake and shag were estimated from their laying dates by adding constants of 27 and 30 days respectively for the incubation periods of these species (https://www.bto.org/understanding-birds/birdfacts). The duration of the chick-rearing period was only available annually with precision in guillemot and razorbill in these populations, so we used a species-specific constant available from the literature to ensure a consistent approach was used across the five species. Fledging dates were therefore calculated by adding average fledging periods to hatching dates (midpoint between minimum and maximum fledging period obtained from the British Trust for Ornithology website: puffin 39 days, kittiwake 43 days, guillemot 21 days, razorbill 19 days, and shag 53 days; https://www.bto.org/understanding-birds/birdfacts; accessed 02/03/2021)

Table S2. Seabird data used to estimate breeding phenology (number of pairs monitored by year), breeding success (number of pairs monitored by year) and number of age 0 sandeel measured from puffin diet samples used in the sandeel Gompertz Growth Model.

Text S3. Sandeel growth model

Model specifications

For individual *i*, *obsTLit* (in mm) was the total length measured at time *t* (day of the year). Growth was modelled using a Gompertz Growth Function (GGF, Fig. S1) to describe *TLit*. An error term accounting for both potential measurement error and deviation from the GGF was incorporated to relate the $obsTL_{it}$ to TL_{it} , similar to the formulation in Reinke et al. (2020). A normal error structure was used with:

$$
obsTL_{it} \sim Normal\left(TL_{it}, \sigma^2_{sampler}\right) \tag{1}
$$

where σ^2 _{sampler} is a gear-specific variance. Gear-specific standard deviations (accounting for sampling error and lack of fit of the GGF with the observed sizes) were small for gears that target larvae (Methot net, Gulf III sampler, and 1m-ringnet) and for dredges as well as puffin chick diet. Due to the relatively small number of observations, standard deviations for grabs and pelagic trawls were much larger. The GGF formulation adopted here uses four parameters. *L0* is an offset value corresponding to size at hatching, *Linfi* is the asymptotic size, K_i is a growth rate coefficient and *Ti* is the date (day of the year) of the inflection point used as a proxy for settlement. The latter two parameters are affected by asynchrony of sandeel with their copepod prey, and therefore, allow the GGF model to incorporate the influence of age 0 sandeel trophic interactions with their prey (Régnier et al. 2017, 2019). The GGF is formulated as:

$$
TL_{it} = L0 + Linf_i \times exp(-exp(-K_i \times (t_i - Ti_i)))
$$
\n(2)

The parameters *Tii*, *Linfi* and *Ki* were estimated at the individual level while initial size, *L0*, was assumed to be constant. We assumed a normal distribution restricted to be positive for parameters *Linfi* and *Tii* with:

$$
Linf_i \sim Normal\left(\mu_{Linf} + y_{Linf,par}, \sigma^2_{Linf}\right) \tag{3}
$$

$$
Ti_i \sim Normal\left(\mu_{Ti} + y_{Tigear}, \sigma^2_{Ti}\right) \tag{4}
$$

A logit-normal variation was assumed for the growth rate K_i , with:

$$
logit(K_i) \sim Normal(\mu_K + y_{K_{year}}, \sigma^2_K)
$$
 (5)

Weakly informative priors were used for μ_{Linf} and μ_{Ti} , with $\mu_{\text{Linf}} \sim \text{Normal}(77,100)$ centred on the average length of age 0 sandeel observed in winter dredge surveys and μ_{Ti} ~Normal(130,100), centred on approximate date at settlement (Gibb et al. 2017, Régnier et al. 2017). A semi-informative prior was used for *L0* with $\mu_{L0} \sim \text{Normal}(6,2)$ based on observed length at hatching in *A. marinus* (Régnier et al. 2018). For μ_K , a vague prior was used with μ_K ~Normal(0,1000). Priors for year-specific parameters *yLinfyear* and *yTiyear* were assigned to a normal distribution centred on 0 with a large variance (Normal(0,1000) while *yKyear* was given a semi informative prior with $yK_{year} \sim Normal(0,0.1)$). For parameters $\sigma_{sampler}$ and σ_{Ti} , tdistributed priors were used with $\sigma \sim t(0,0.0004,3)$ restricted to be positive. For σ_{Linf} and σ_K , truncated positive Normal distributions were used (means 10 and 0.1, variances 0.25 and 0.001 respectively). A total of 300,000 iterations with an adaptive phase of 50,000 iterations, a burnin of 50,000 iterations and 200,000 iterations were used to estimate the posterior distributions of the parameters with 2 parallel chains. Convergence was assessed visually through trace-plots and analytically through the potential scale reduction factor (PSRF, Gelman & Rubin 1992) for which an upper 0.975 quantile of ≤ 1.2 has been given as a rule of thumb to indicate convergence (Smith 2007). The program JAGS (Plummer 2003) was used to run the model through an R interface (R 4.1.2, R Core Team 2021) using the *rjags* package. Goodness of fit was assessed visually for each year plotting observations, annual mean and individual predictions.

Model outputs

Convergence was achieved for all model parameters presented in Fig. 2 in the main text, and the model provided a good fit to the data overall (Fig. S2). Sampling processes unaccounted for in the model led to localised lacks of fit (e.g. year 2010, Figs. S2-S3) but had minor effects on the predictive power of the model. In particular, age 0 sandeel lengths obtained from puffin chick diet showed a decrease throughout July (starting around day 180 where day $1 = 1$ Jan) in a number of years, leading to estimated sizes being larger than observations for these periods (Fig. S3). This apparent lack of large age 0 sandeel toward the end of the chick rearing period in puffin is consistent with a condition-dependent decrease in sandeel activity and coincided with the time age 0 sandeel reached their asymptotic size during their first year (Fig. S3). This observation supports the use of the date at which age 0 sandeel reached 95% of their asymptotic size as a proxy for the end of the period of age 0 availability.

The main parameters of interest (*K*, *Linf* and *Ti*) were only weakly correlated with all correlation coefficients in the range -0.06 to 0.29. Settlement date (*Ti*) was positively related to sandeel hatching date ($\mathbb{R}^2 = 0.57$, $F_{1,15} = 19.78$, p=0.0005, Fig. S4), such that later hatching translated into later settlement. The growth rate parameter (*K*) increased with average sea temperature in the 60 days following sandeel hatching and with the measure of synchrony

between sandeel and their copepod prey (Régnier et al. 2019) but decreased with their interaction (Fig. S5). In particular, the positive effect of temperature on growth rate was associated with good synchrony between sandeel and their prey, while the cost of trophic asynchrony on growth rate became apparent at high temperatures (Fig. S5). Estimates of *sandeelSTART* and *sandeel END* were related to the estimated growth rate and date of settlement (Fig. S6).

Our ability to identify trophic linkages, which propagated from lowest trophic levels to seabird predators, depended on an accurate description of age 0 sandeel (*A. marinus*) growth. The Gompertz Growth Model (GGM) used multiple data sources allowing a good coverage of the juvenile period and was validated by a good fit of the individual growth trajectories to the data. This model enabled the estimation of key parameters of age 0 sandeel growth and phenology, necessary for the study of trophic mismatch with their avian predators. The twofold variation of the estimated age 0 sandeel growth in the 17-year time-series was best explained by an interaction between trophic asynchrony with their copepod prey and average temperature in the 2 months following hatching. As the trophic asynchrony between age 0 sandeel and their copepod prey is itself indirectly related to temperature (Régnier et al. 2019), age 0 sandeel growth is therefore temperature sensitive and a parameter of interest in the context of climate change. In terms of phenology, the date of settlement and the date at which sandeel reach 50 mm and appeared in seabird diet could be estimated and marked the start of the period of age 0 sandeel availability. Estimated dates of settlement were within the range of settlement dates back-calculated from age 0 otoliths (Gibb et al. 2017). Year-specific values from the two methods were positively correlated confirming that the date of the inflection point in the growth model was an appropriate proxy for settlement date. There was also a strong relationship between this parameter, hatching dates estimated from otoliths and the appearance of larvae (Régnier et al. 2017, 2019), providing further confidence in the use of these parameters. As our measure of age 0 sandeel availability integrates prey quality by reflecting the availability of age 0 sandeel of a size > 50 mm, some years were characterised by a very short period of age 0 sandeel availability (e.g. 2007, Fig. 2a in the main text). For such years, while age 0 sandeel were present in seabird diets, the fish from puffins were very small and the corresponding energetic content and profitability would have been very low (Hislop et al. 1991). The assumption that, as in older sandeel, availability of age 0 sandeel decreased as a trade-off between growth opportunity and predation risk (Haynes et al. 2007, Bergstad et al. 2002) was supported by the data. A decrease in the size range of age 0 sandeel caught by puffins coincided with the date age 0 sandeel reached their maximum size during the first summer (Fig. S3). Both the estimated dates of the start and end of the period of age 0 sandeel availability were defined by variability in the date of settlement and growth rate (Fig. S6). As growth in ectotherms is sensitive to temperature, age 0 sandeel phenology and the degree of synchrony between age 0 sandeel availability and seabird breeding phenology are therefore likely influenced by climate change. Fig S7 shows the measure of overlap between age 0 sandeel availability and seabird chick-rearing period was related to both the date of the start and end of the period of age 0 sandeel availability.

Fig. S1. Directed Acyclic Graph (DAG) corresponding to the sandeel growth model. Squares denote constants, circles denote variables. Solid arrows denote stochastic processes while dotted arrows denote deterministic processes.

Fig. S2. Estimated trajectories for age 0 sandeel for years 2000 to 2016. Individual data are indicated in black, solid red lines correspond to the annual trajectories and solid blue lines correspond to individual trajectories for a sample of 150 individuals each year.

Fig. S3. Deviations between sizes of age 0 sandeel captured by puffins and predictions from the Age 0 growth model (grey points, left axis). The density distribution of the end of sandeel availability is indicated by a thick line (black, right axis).

Fig. S4. Relationship between settlement date (parameter Ti) and median hatch day estimated for sandeel from Régnier et al. (2019). The mean response is indicated by a solid line and the 95% confidence interval is shaded.

Fig. S5. Relationship between average sea temperature (°C), estimated growth rate [K] and the degree of trophic mismatch between *Ammodytes marinus* and its copepod prey (absolute value).

Fig. S6. Relationship between (a) the start of the period of age 0 sandeel availability and the estimated growth rate [K] and date of settlement [Ti] and (b) the end of the period of age 0 sandeel availability and the estimated growth rate [K] and date of settlement [Ti].

Fig. S7. Relationship between the overlap between the chick rearing period (overlap index) and the date of the start and end of the period of age 0 sandeel availability for the five seabird species considered.

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